

Supplementary Online Content 4

The Multicenter Uveitis Steroid Treatment (MUST) Trial and Follow-up Study Research Group. Association between long-lasting intravitreal fluocinolone acetonide implant vs systemic anti-inflammatory therapy and visual acuity at 7 years among patients with intermediate, posterior or panuveitis. *JAMA*. doi: 10.1001/jama.2017.5103

Statistical Analysis Plan for MUST Follow-up Study

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**Association Between Long-lasting Fluocinolone
Acetonide Intravitreal Implant vs Systemic
Anti-inflammatory Therapy and Visual Outcomes at
Seven Years Among Patients With Intermediate,
Posterior, or Panuveitis**

Statistical Analysis Plan

Version 3
19 March 2017

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4 The primary outcome measurement in the MUST Trial Follow-up Study is the change in BCVA from
5 enrollment to 7 years of follow-up. The statistical methods used to analyze the long-term follow-up
6 in the continuation study will be the similar to those outlined in the original MUST Trial. For the
7 MUST Follow-up Study, an observational study, the primary analysis still will be according to the
8 original randomization group, following the principles of Intent to Treat (ITT) in that all patients will
9 be grouped according to their assigned therapy regardless of the actual treatment received. This
10 allows us to answer the real-world question of the long-term prognosis of an individual undergoing
11 an intended course of therapy. Secondary analyses will be based on the treatment received, and
12 the ITT analysis will be used as a point of departure from which a causal analysis based in time-
13 varying covariates and propensity scores to account for selection and dropout effects will be built.
14 The relation between the ITT analysis and the causal modeling will be of interest in its own right, but
15 our principal focus will be on understanding the long-term trajectories via the analysis of the
16 observational data. One of the treatment-based analyses will focus upon those individuals who
17 receive treatment according to the protocol specifications. Individuals (or eyes) will be included in
18 this analysis if they received their assigned treatment within six months of randomization. For
19 individuals (or eyes) assigned to systemic treatment, follow-up will be censored at time of
20 implantation, if an implant is placed. All primary analyses will be replicated by at least two different
21 analysts.

22

23 As during the trial phase of MUST, evaluation of continuous outcomes over time (such as change in
24 BCVA) and binary outcomes over time (such as normal/abnormal IOP) will use a repeated measures
25 analysis with Gaussian or logit links and accounting for the nested correlations between
26 observations over time and (when applicable) between eyes of the same participant.¹ Both absolute
27 (percent) and relative (odds ratios) comparisons will be presented for binary outcomes using
28 unadjusted and adjusted logistic models, respectively. The primary analysis will be based upon
29 generalized estimating equations (GEE) since our goal is to summarize the marginal (i.e. population
30 average) effects. Sensitivity analyses using mixed effects models will also be explored; however,
31 since the goal is the population as opposed to individual level effects, they will primarily be used to
32 assess the impact of missing data, incorporation of clinic (as a random effect), and alternate
33 covariance structures. The mixed effects analysis also will allow for the exploration of the effect of
34 clinic on the results. For short to moderate term follow-up, a saturated mean model, including visit
35 and visit by treatment interaction terms, will be used. An unstructured covariance matrix will be
36 used to model the within-eye repeated measurements augmented by random effects to induce
37 cross-sectional between-eye associations in the clinical trials. However, the unstructured covariance
38 model will not be feasible for the long-term follow-up study, because a large number of parameters
39 need to be estimated. Therefore, we will replace it by a Toeplitz covariance structure or, for the
40 mixed effects model, a structure composed of a first-order, auto-regressive process (an AR(1)
41 model) along with a random intercept. These structures allow for correlation to decrease with
42 increasing time-separation.

43

44 Evaluation of risk factors for time-to-event outcomes such as incidence of ME, cataracts, IOP
45 elevation or glaucoma, as well as time from ME to remission will be performed using Cox

46 proportional hazards regression as well as parametric time-to-failure models, such as gamma
47 models², including a random effects term to account for the between eye correlation.
48 Implementation of these models allows for clustering (within participants and within eyes),
49 assessment of recurrent events, and incorporation of time-dependent covariates. When the
50 relevant follow-up time in the analysis reflects the clinical time scale (e.g. time since diagnosis of
51 uveitis), it is necessary to incorporate the prevalent cases (longstanding diagnosis of uveitis) into the
52 analysis. To accommodate the prevalent cases we will use the staggered entry technique³ which is
53 one method of adjusting for potential survival bias.

54 The analysis compares the event rates among participants with similar duration of disease and then
55 combines over these comparisons. Event rates for multiple recurring events, e.g. the number of
56 adverse events, will be modeled using Poisson regression or Negative Binomial regression, including
57 a random effects term to account for the between eye correlation.²
58

59 Cross-sectional comparisons between subgroups (e.g. those who did and did not complete follow-
60 up) will be performed. For person-level characteristics, χ^2 tests or Fisher's exact test and Wilcoxon
61 rank-sum tests will be used to compare categorical and continuous variables, respectively. For eye-
62 level characteristics, mixed effects models will be used.

63

64 All analyses will be performed both unadjusted and adjusted for potential confounders. Effect
65 modification due to factors such as disease location, systemic disease, gender and race also will be
66 explored when appropriate. Robust standard errors will be computed using statistical program-
67 based approaches when available and a bootstrap with the individual as the sampling unit, when a
68 pre-programmed approach is not available. The bootstrap was primarily used to adjust for
69 between-eye correlation for ocular outcomes.

70

71 The bootstrap routine for estimating standard errors, confidence intervals, and p-values is as
72 follows. First, the original estimates will be obtained by analyzing the trial and follow-up study data
73 adjusting for longitudinal within-eye measurements assuming a working independence structure for
74 the between-eye correlations, i.e. treating each eye as an independent observation. A total of 5000
75 iterations will be used to calculate the bootstrap estimates. The sampling unit is the individual (as
76 opposed to the eye) to preserve the between-eye correlation structure. Within each subgroup
77 defined by treatment and disease laterality (unilateral systemic, bilateral systemic, unilateral
78 implant, and bilateral implant), we will sample with replacement to create a new cohort of the same
79 size. The resulting datasets will maintain sample size, both in terms of individuals and eyes, found
80 within each treatment group in the original dataset. We will then apply the model developed for
81 the original dataset to each of the bootstrap samples. The standard deviation of the bootstrap
82 estimates will be used as the standard error for the original estimates. Assuming a normal
83 distribution, t-test statistics, p-values, and 95% confidence intervals may be constructed around the
84 original parameter estimates using the bootstrap standard error. Alternately, the 95% confidence
85 interval may be constructed by extracting the 2.5- and 97.5-percentiles from the bootstrap
86 estimates.

87

88 The follow-up study focuses on a set of primary research questions and related analyses. All tests
89 will be considered statistically significant at the 0.05 level. However, a large number of comparisons
90 are planned for secondary outcomes and caution is needed in the reporting of interpretation of
91 these results. As recommended by Wang et al.⁴ our primary focus for these outcomes will be on the
92 parameter estimates and confidence intervals rather than p-values. Several methods of adjusting p-
93 values for multiple comparisons exist, however no clear consensus as to the most appropriate
94 method is available and it is difficult if not impossible to quantify the number of comparisons. In
95 general, issuing cautions is sufficient, but for identifiable and related sets of estimates we will do
96 adjustments. We expect that related sets of estimates will have a high positive correlation, making a
97 Bonferroni correction extremely conservative. Therefore, we will also explore using an estimate of
98 the covariance matrix for these related sets based upon a bootstrap approach and also estimate the
99 null distribution of the minimum p-values for the multivariate distribution of Z-scores using a global
100 null hypothesis permutation distribution or the multivariate normal cumulative distribution program
101 is R. A variety of sensitivity analyses will be performed in order to determine the potential for bias
102 due to modeling assumptions, missing data, and potential biases (especially for epidemiologic
103 analyses). Transformations (e.g., log) of continuous outcomes will be considered when violations of
104 the Gaussian assumption occur.

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106 Single imputation techniques including ‘best’ and ‘worst’ case scenarios will be used to identify the
107 magnitude of the potential impact of missing data. For example, we will calculate the magnitude of
108 the effect sizes that would need to be observed among those individuals with missing data in order
109 to change our inference or conclusions. Likelihood based methods such as mixed effects models
110 that are robust to data that is missing at random will be explored.⁵ In addition, multiple imputation
111 (MI) and pattern mixture approaches will be used to assess the impact of missing data.⁵⁻⁶
112 Specifically, we will retain all features of the primary analysis other than how missing values for
113 visits beyond the last one with a measured value are handled. As is implicit in the primary analysis
114 where missing data indicators are used (e.g. “.” in SAS, “NA” in R), we will treat all missing values,
115 before the last measured one, as Missing at Random (MAR). We will impute other missing values to
116 generate 10 pseudo-complete records for each such individual by sampling from a joint predictive
117 distribution for the missing data given the observed data. We will use a covariance matrix equal to
118 the estimated covariance matrix from the primary analysis, but will “take control” of the prediction
119 mean. Varying the mean of the predictive distribution allows us to assess the sensitivity of our
120 results to a variety of missing data scenarios. The pattern mixture model approach stratifies
121 participants on their pattern of missing data, estimates stratum-specific parameters and then
122 combines estimates over strata using inverse variance weights. The approach is similar to stratified
123 analysis to adjust for potential confounders and allows for comparison of stratum-specific estimates.
124 We stratify by three patterns: complete data, at least one internal measurement missing, and
125 closeout weight missing (along with any other missingness).

126

127 The presence of time-varying confounders that are themselves affected by prior levels of exposure
128 may produce large biases in the estimation of causal effects using standard statistical analyses. A
129 classic example, demonstrated by Cole et al.⁷ shows that adjusting for time-varying CD4+ T cell
130 count can dramatically reduce the estimated benefit of HAART on time to AIDS or death. One
131 potential example in the MUST Trial Follow-up Study would be adjusting for cataract status when
132 estimating the time to vision loss for participants receiving regional corticosteroid treatments, since
133 these treatments are known to cause cataracts. Another example influencing participant based
134 outcomes is utilization of immunological agents with systemic corticosteroid treatments. Marginal
135 structural models can be used to correct both for bias induced by a variable affected by exposure

136 and for bias induced by loss to follow-up⁸ The technique employs time-varying inverse probability
137 weights in the place of standard covariate adjustments. We will explore the presence of time-
138 varying confounding by comparing standard Cox proportional hazards models with marginal
139 structural models.
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142 **Document revision history**

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145 **Version 1 of this statistical analyses plan was developed from the MUST Trial (Version 2.4) and**
146 **Follow-up Study (Version 4.4) Protocols and was submitted as part of the original JAMA**
147 **submission. Substantive clarifications to the original statistical analysis plan were made in**
148 **response to reviewer queries in Version 2 (1 March 2017). All revisions are captured in the**
149 **tracked change versions of the documents archived at the Coordinating Center.**

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151 **1. Analysis datasets (p.2)**

152 Clarification – The primary analysis will be based upon the principles of intention to treat (ITT)

153 **2. Analytic models (p.2)**

154 Clarification – Generalized estimating equations will be used for the primary analyses with
155 secondary analyses based upon mixed effects models.

156 Clarification – Specified that both absolute and relative effects would be calculated for binary
157 outcomes in response to the JAMA review. Additional details of the absolute analysis are
158 provided on p. 3.

159 **3. Bootstrap analysis (p. 3)**

160 Clarification – Additional details on the bootstrap methods to adjust for between-eye correlation
161 and to estimate absolute effects for binary outcomes were added at the request of the JAMA
162 review.

163 **4. Missing data (p. 3)**

164 Clarification – Additional details of the missing data analyses including ‘best’ and ‘worst’ case
165 scenarios as well as mixed effects models were added.

166 **5. References (p.6)**

167 Clarification – References were corrected and a reference list added to the file.

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Version 3 (18 March 2017). incorporates additional clarifications in response to the second round of review All revisions are captured in the tracked change versions of the documents archived at the Coordinating Center.

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175 **1. Treatment-based analyses (p.2)**

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Clarification – Specifics of the criteria for the per-protocol treatment-based analysis were added.

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178 **2. Role of mixed effects model (p.2)**

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Clarification – Exploration of the effect of clinic will be performed using mixed effects models

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181 **3. Cross-sectional comparisons of characteristics (p.3)**

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Addition – Specifics for cross-sectional comparisons are specified.

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184 **4. Level of statistical significant (p.3)**

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Clarification – The statistical level of significance (0.05) was added to the discussion of multiple comparisons.

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193 References

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